5010 solutions, Assignment 8. Chapter 5: 1, 3, 6, 9, 10, 12, 18, 21, 24.

1. $f(x,y) = \binom{2}{x,y,2-x-y} (1/6)^x (1/6)^y (4/6)^{2-x-y} = \binom{2}{x,y,2-x-y} 4^{2-x-y}/36$ for $x \ge 0, y \ge 0$, and $x + y \le 2$. We find that f(0,0) = 16/36, f(0,1) = f(1,0) = 8/36, f(1,1) = 2/36, and f(0,2) = f(2,0) = 1/36. The marginals are both binomial (2,1/6) so E[X] = E[Y] = 1/3 and Var(X) = Var(Y) = 2(1/6)(5/6) = 5/18. Next E[XY] = 0(34/36) + 1(2/36) = 1/18, so $Cov(X,Y) = 1/18 - (1/3)^2 = -1/18$ and $\rho(X,Y) = (-1/18)/(5/18) = -1/5$.

3. Let X and Y be the results of two dice throws. The X + Y and X - Y are uncorrelated but not independent. Indeed, $Cov(X + Y, X - Y) = E[(X + Y)(X - Y)] = E[X^2] - E[XY] + E[YX] - E[Y^2] = E[X^2] - E[Y^2] = 0$ and, if X + Y = 2, then X = Y = 1, so X - Y = 0. The example in the book's answers appears wrong because Y is constant, and a constant is independent of any random variable.

6. The values of Y are $0, 1, 0, -1, 0, 1, 0, -1, \ldots$, each with probability 1/(4n). The values of Z are $1, 0, -1, 0, 1, 0, -1, 0, \ldots$, each with probability 1/(4n). The joint probability mass function of Y and Z is f(0, 1) = 1/4, f(1, 0) = 1/4, f(0, -1) = 1/4, and f(-1, 0) = 1/4. Therefore Y + Z is ± 1 with probability 1/2 each. Orthogonality means E[YZ] = 0 and this expectation is 0 because YZ is the 0 random variable.

9. (a) $P(U = n) = P(X_1 = 1)P(U = n | X_1 = 1) + P(X_1 = -1)P(U = n | X_1 = -1) = pp^{n-1}q + qq^{n-1}p$, which is p times a geometric(q) distribution plus q times a geometric(p) distribution. So E[U] = p(1/q) + q(1/p). $P(V = n) = P(X_1 = 1)P(V = n | X_1 = 1) + P(X_1 = -1)P(V = n | X_1 = -1) = pq^{n-1}p + qp^{n-1}q$, which is p times a geometric(p) distribution plus q times a geometric(q) distribution. So E[V] = p(1/p) + q(1/q) = 2.

(b) $P((U,V) = (m,n)) = P(X_1 = 1)P((U,V) = (m,n) | X_1 = 1) + P(X_1 = -1)P((U,V) = (m,n) | X_1 = -1) = pp^{m-1}q^np + qq^{m-1}p^nq = p^{m+1}q^n + q^{m+1}p^n$ for $m, n \ge 1$. Next,

$$\begin{split} E[UV] &= p \sum_{m,n \ge 1} mn \cdot p^{m-1}q \cdot q^{n-1}p + q \sum_{m,n \ge 1} mn \cdot q^{m-1}p \cdot p^{n-1}q \\ &= pE[\text{geometric}(p)]E[\text{geometric}(q)] + qE[\text{geometric}(q)]E[\text{geometric}(p)] \\ &= \frac{p}{pq} + \frac{q}{qp} = \frac{1}{p} + \frac{1}{q}. \end{split}$$

Hence

$$\operatorname{Cov}(U,V) = \frac{1}{p} + \frac{1}{q} - 2\left(\frac{p}{q} + \frac{q}{p}\right) = \frac{1 - 2(p^2 + q^2)}{pq} = \frac{4pq - 1}{pq}$$

The correlation requires finding the variances of U and V, which can be done by the same methods as above.

10. P((X,Y) = (i,j)) = 1/[n(n-1)] for all $i \neq j$. So the joint distribution is uniform over the pairs (i,j) with i, j = 1, ..., n and $i \neq j$. It follows that (or

it is obvious, à priori) that X and Y are uniform over 1, 2, ..., n and $E[U] = E[V] = (1 + 2 + \dots + n)/n = (n + 1)/2$. So

$$\begin{split} E[XY] &= \frac{1}{n(n-1)} \sum_{i \neq j} ij = \frac{1}{n(n-1)} \left(\sum_{i,j} ij - \sum_{i} i^2 \right) \\ &= \frac{1}{n(n-1)} \left(\left(\frac{n(n+1)}{2} \right)^2 - \frac{n(n+1)(2n+1)}{6} \right) = \frac{(n+1)(3n+2)}{12} \end{split}$$

hence

$$\operatorname{Cov}(X,Y) = \frac{(n+1)(3n+2)}{12} - \frac{(n+1)^2}{4} = -\frac{n+1}{12}$$

Next, $E[U^2] = n(n+1)(2n+1)/(6n) = (n+1)(2n+1)/6$, so $Var(U) = (n+1)(2n+1)/6 - (n+1)^2/4 = (n-1)(n+1)/12$. Hence $\rho(X,Y) = -1/(n-1)$. This tends to 0 as $n \to \infty$.

- 12. (a) f(i,j) = 1/36 if i < j, f(i,i) = (6-i+1)/36, f(i,j) = 0 if i > j.
- (b) f(i,j) = 2/36 if i < j, f(i,i) = 1/36, f(i,j) = 0 if i > j. (done in class) (c) f(i,j,k) = 1/36 if $k = \max(i,j)$, f(i,j,k) = 0 otherwise.

For Cov(U, V), we need E[UV]. There is a trick to get this one: Just note that UV = XY and so $E[UV] = E[XY] = E[X]E[Y] = (7/2)^2$. Next $E[U] = (11 \cdot 1 + 9 \cdot 2 + 7 \cdot 3 + 5 \cdot 4 + 3 \cdot 5 + 1 \cdot 6)/36 = 91/36$ and $E[V] = (1 \cdot 1 + 3 \cdot 2 + 5 \cdot 3 + 7 \cdot 4 + 9 \cdot 5 + 11 \cdot 6)/36 = 161/36$. So $\text{Cov}(U, V) = 49/4 - (91/36)(161/36) = (36 \cdot 9 \cdot 49 - 91 \cdot 161)/6^4 = 1225/6^4 = 5^27^2/6^4$. Finally,

$$E[XYV] = \frac{1}{36} \sum_{i=1}^{6} \sum_{j=1}^{6} ij \max(i, j)$$

This would be a bit tedious to compute by hand (and unreliable), so

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FOR i=1 to 6
   FOR j=1 to 6
    LET sum=sum+i*j*max(i,j)
   NEXT j
NEXT i
PRINT sum
END
  2275
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The answer is $2275/36 = 91(5/6)^2 = 63.19\overline{4}$.

18. (a)

$$E[Z] = E[\min(c, X)] = \sum_{n=1}^{\infty} P(\min(c, X) \ge n) = \sum_{n=1}^{c} P(X \ge n)$$
$$= \sum_{n=1}^{c} (1 - p_1)^{n-1} = \frac{1 - (1 - p_1)^c}{p_1}.$$

Book's answer is wrong.

(b) $P(\min(X,Y) \ge n) = P(X \ge n, Y \ge n) = P(X \ge n)P(Y \ge n) = (1-p_1)^{n-1}(1-p_2)^{n-1} = [(1-p_1)(1-p_2)]^{n-1}$, and this implies that $\min(X,Y)$ has a geometric distribution with parameter $1-(1-p_1)(1-p_2) = p_1+p_2-p_1p_2$, hence its mean is the reciprocal of this, $1/(p_1+p_2-p_1p_2)$. Book's answer is wrong.

21. (a) $f_{X+Y}(z) = \sum_{x} f_X(x) f_Y(z-x) = (n+1)^{-2}(z+1)$ for z = 0, 1, ..., nand $f_{X+Y}(z) = (n+1)^{-2}(n-(z-n)) = (n+1)^{-1}(2n-z+1)$ for z = n, n+1, ..., 2n.

(b) $f_{X-Y}(z) = \sum_{x} f_X(z+x) f_Y(x) = (n+1)^{-2} (\min(n, n-z) - \max(0, -z) + 1)$ for $z = -n, -n+1, \ldots, n-1, n$. For $z = 0, 1, \ldots, n$, this gives $f_{X-Y}(z) = (n+1)^{-2}(n-z+1)$. For $z = -n, -n+1, \ldots, 0$, this gives $f_{X-Y}(z) = (n+1)^{-2}(n+z+1)$.

24.

$$P(X+Y=n) = \sum_{m} P(X=m, Y=n-m) = \sum_{m=0}^{n} (e^{-\lambda} \lambda^{m}/m!)(e^{-\mu} \mu^{n-m}/(n-m)!)$$
$$= e^{-(\lambda+\mu)} \left[\sum_{m=0}^{n} \binom{n}{m} \lambda^{m} \mu^{n-m}\right]/n! = e^{-(\lambda+\mu)} (\lambda+\mu)^{n}/n!.$$

$$P(X = k \mid Z = n) = \frac{P(X = k, Z = n)}{P(Z = n)} = \frac{P(X = k, Y = n - k)}{P(Z = n)}$$
$$= \frac{(e^{-\lambda}\lambda^k/k!)(e^{-\mu}\mu^{n-k}/(n-k)!)}{e^{-(\lambda+\mu)}(\lambda+\mu)^n/n!} = \binom{n}{k}p^k(1-p)^{n-k}$$

where $p = \lambda/(\lambda + \mu)$.